NMAI059 Probability and statistics 1 Class 13

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Statistics - What have we learnt

- ▶ basic setup: we consider random sample X₁,..., X_n from distribution F_ϑ describes the measurement process, all ways how could it go
- we measure data particular numbers x₁,..., x_n, so called realization of random sample, — what did we really measure
- 1. point estimation: determine best possible number, estimate for the parameter ϑ , or some function of it, $g(\vartheta)$.
- 2. interval estimation: determine an interval, that contains the unknown parameter ϑ with a large probability

3. hypothesis testing

Overview

Hypothesis testing

Goodness of fit test

Linear regression

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Hypothesis testing - illustration

- We want to test, if a coin is fair.
- H_0 : it is fair
- H₁: not fair ("Scientists discovered, that casino XY uses loaded coin.")
- Results: Reject H_0 /don't reject H_0
- ► Type I error: false rejection. We reject *H*₀, even if it is true. Embarassing.
- Type II error: false non-rejection. We don't reject H₀, even if it is false. Unused opportunity.
- ▶ Need to find k such that we will reject H_0 if |S n/2| > k.

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Hypothesis testing – general approach

- We choose an appropriate statistical model.
- We choose *significance level* α : prob. of false rejection of H_0 . Typically $\alpha = 0.05$ (medicine/psychology much less in high-energhy physics).
- We determine *test statistics* $S = h(X_1, ..., X_n)$, that we will determine from the measured data.
- ► We determine *rejection region* set *W*.
- We measure x_1, \ldots, x_n so-called realizations of X_1, \ldots, X_n .
- Decision rule: we reject H_0 iff $h(x_1, \ldots, x_n) \in W$.

$$\qquad \qquad \bullet \quad \alpha = P(h(X) \in W; H_0)$$

- $\beta = P(h(X) \notin W; H_1) \dots$ strength of the test
- often we do not choose α in advance but compute so-called *p*-value: minimal α, for which we would reject H₀.

Hypothesis testing – an example

• X_1, \ldots, X_n random sample from $N(\vartheta, \sigma^2)$

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- $\blacktriangleright \sigma^2$ known
- $\blacktriangleright H_0: \vartheta = 0 \qquad H_1: \vartheta \neq 0$

Hypothesis testing – an example

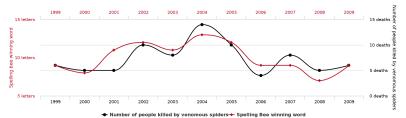
• X_1, \ldots, X_{n_1} random sample from $Ber(\vartheta_X)$

- Y_1, \ldots, Y_{n_2} random sample from $Ber(\vartheta_Y)$
- $\blacktriangleright H_0: \vartheta_X = \vartheta_Y \qquad H_1: \vartheta_X \neq \vartheta_Y$

p-hacking

- we first gain data, then look for interesting stuff
- given enough data, there will be random coincidences
- even worse, we may test, until we get the desired outcome
- reproducibility after exploratory analysis of the data we make an independent measurement and a confirmatory analysis.
- or we split the data in advance to a part for hypothesis formation and part for verification ... simple example of cross validation

Letters in winning word of Scripps National Spelling Bee



Number of people killed by venomous spiders

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χ_k^2 – chi-square distribution

Definition $Z_1, \ldots, Z_k \sim N(0, 1)$ *i.i.d.* The distribution of r.v.

$$Q = Z_1^2 + \dots + Z_k^2$$

is called chi-square(d) with k degrees of freedom (really k!), and denoted $\chi^2_k.$

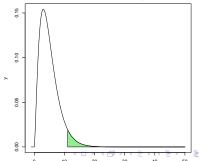
$$\blacktriangleright \mathbb{E}(Q) = k$$
 (easy)

• var(Q) = 2k (fyi, you don't have to remember this)

 density can be written by a reasonable formula

•
$$Q \doteq N(k, 2k)$$

for large k (CLT)



Multinomial and categorical distribution

Definition

Given $p_1, \ldots, p_k \ge 0$ so, that $p_1 + p_2 + \cdots + p_k = 1$. we repeat *n*-times an experiment with *k* possible outcomes, where the *i*th has probability p_i

 $X_i :=$ how many times we got the *i*-th outcome (X_1, \ldots, X_k) has multinomial distribution with parameters $n, (p_1, \ldots, p_k)$.

- trivial example: X_i = number of die rolls that equaled i
- important example: X_i = number of occurences of *i*-th letter,

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•
$$P(X_1 = x_1, \dots, X_k = x_k) = \binom{n}{x_1, \dots, x_k} p_1^{x_1} \dots p_k^{x_k}$$

Pearson χ^2 statistics

► (X₁,...,X_k) – multinomial distribution with parameters n, (p₁,...,p_k) as above

$$\triangleright E_i := \mathbb{E}(X_i) = np_i$$

• Pearson χ^2 statistics is the function

$$\chi^2 = T := \sum_{i=1}^k \frac{(X_i - E_i)^2}{E_i}$$

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• Theorem $T \xrightarrow{d} \chi^2_{k-1}$

Goodness of fit test

- ► (X₁,...,X_k) multinomial distribution with parameters n, (ϑ₁,...,ϑ_k) as above
- \blacktriangleright *n* known, ϑ unknown
- Null hypothesis $H_0: \vartheta = \vartheta^*$ for some given ϑ^*
- $E_i := n\vartheta_i^*$ for all i
- We use the statistics $\chi^2 = T := \sum_{i=1}^k \frac{(X_i E_i)^2}{E_i}$
- We reject H_0 iff $T > \gamma$
- $\blacktriangleright \ \gamma := F_Q^{-1}(1-\alpha), \text{ where } Q \sim \chi^2_{k-1}$
- $\blacktriangleright \ P(\mathsf{I type error}) = P(T > \gamma; H_0) \rightarrow P(Q > \gamma) = \alpha$

Goodness of fit test - example

We roll a die repeatedly (600 times). The numbers 1 upto 6 came up with frequencies 92, 120, 88, 98, 95, 107.

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Is the die fair?

Extensions

- ► To study a distribution of an arbitrary r.v. *Y* we can pick "bins" B_1, \ldots, B_k (a partition of \mathbb{R}) and look how often $Y \in B_i$ (this will be measured by r.v. X_i).
- Similar test for independence of discrete random variables.

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Linear regression – the problem

• data:
$$(x_i, y_i)$$
 for $i = 1, \ldots, n$

• goal:
$$y = \vartheta_0 + \vartheta_1 x$$

we measure how good fit we have by the quadratic error:

$$\sum_{i=1}^{n} (y_i - (\vartheta_0 + \vartheta_1 x_i))^2$$

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Linear regression – solution

► To minimalize

$$\sum_{i=1}^{n} (y_i - (\vartheta_0 + \vartheta_1 x_i))^2$$

the optimal parameters are

$$\hat{\vartheta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}, \qquad \hat{\vartheta}_0 = \bar{y} - \vartheta_1 \bar{x},$$

where
$$\bar{x} := (x_1 + \dots + x_n)/n$$
, $\bar{y} := (y_1 + \dots + y_n)/n$.

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Linear regression – why sum of squares?

We assume that x₁,..., x_n are fixed, y_i is a realization of a r.v.

$$Y_i = \vartheta_0 + \vartheta_1 x_i + W_i$$

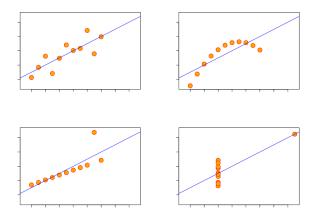
- $W_i \sim N(0, \sigma^2)$ for all $i; W_1, \ldots, W_k$ iid
- maximal likelihood:

$$L(y;\vartheta) = \prod_{i=1}^{n} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(y_i - \vartheta_0 - \vartheta_1 x_i)^2}{2\sigma^2}}$$

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$$\blacktriangleright \ \ell(y;\vartheta) = \log L(y;\vartheta) = a + b \sum_{i=1}^{n} (y_i - \vartheta_0 - \vartheta_1 x_i)^2$$

Limits of regression

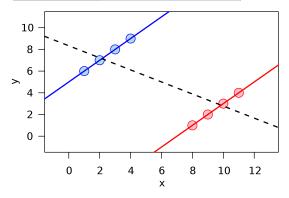


(data: Francis Anscombe 1973, image: wikieditor Schutz)

- nonlinear regression
- logistic regression

Simpson's paradox

| Treatment Stone size | Treatment A | Treatment B |
|-------------------------|--------------------------|----------------------------|
| Small stones | Group 1 93% (81/87) | Group 2 87% (234/270) |
| Large stones | Group 3 73% (192/263) | <i>Group 4</i> 69% (55/80) |
| Both | 78% (273/350) | 83% (289/350) |



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