

Ex. "WASTE" $\begin{pmatrix} Y_1 \\ X_1 \end{pmatrix}, \dots, \begin{pmatrix} Y_n \\ X_n \end{pmatrix}$ i.i.d.
 BIO WASTE $\swarrow$ $\nwarrow$ TOTAL

$$\theta_x = \frac{E Y_1}{E X_1}, \quad \hat{\theta}_n = \frac{\bar{Y}_n}{\bar{X}_n}$$

PREVIOUSLY:

$$\sqrt{n}(\hat{\theta}_n - \theta_x) \xrightarrow[n \rightarrow \infty]{d} N(0, \sigma^2),$$

$$\text{WHERE } \sigma^2 = \frac{1}{(E X_1)^2} \text{var}(Y_1 - \theta_x X_1)$$

ESTIM. OF σ^2 :

$$\hat{\sigma}_n^2 = \frac{1}{(\bar{X}_n)^2} \frac{1}{n-1} \sum_{i=1}^n (Y_i - \hat{\theta} X_i)^2$$

STAND. ASYMPT. CONF. INT

$$\left(\hat{\theta}_n \pm n^{-1/2} \alpha_{1/2} \frac{\hat{\sigma}_n}{\sqrt{n}} \right)$$

NON PARAM. BOOTSTRAP

$$\begin{pmatrix} X_{1,b}^* \\ Y_{1,b}^* \end{pmatrix}, \dots, \begin{pmatrix} X_{n,b}^* \\ Y_{n,b}^* \end{pmatrix} \text{ i.i.d.}$$

$$\leadsto R_{n,b}^* = \sqrt{n}(\hat{\theta}_{n,b}^* - \hat{\theta}_n), b=1, \dots, B$$

$$\hat{\theta}_{n,b}^* = \frac{\bar{Y}_{n,b}^*}{\bar{X}_{n,b}^*}$$

BASIC BOOT. CONF. INT

$$\left(\hat{\theta}_n - \frac{n_{n,B}^*(1-\alpha/2)}{\sqrt{n}}, \hat{\theta}_n - \frac{n_{n,B}^*(\alpha/2)}{\sqrt{n}} \right),$$

WHERE $n_{n,B}^*(\alpha)$ IS "SAMPLE" α -QUANTILE FROM THE "BOOT. SAMPLE" $R_{n,1}^*, \dots, R_{n,B}^*$

PERCENTILE CONF. INT.

$$\left(q_{n,B}^*(\frac{\alpha}{2}), q_{n,B}^*(1-\frac{\alpha}{2}) \right),$$

WHERE $q_{n,B}^*(\alpha)$ IS THE "SAMPLE" α -QUANTILE
FROM THE "BOOT. SAMPLE": $\hat{\sigma}_{n,1}^*, \dots, \hat{\sigma}_{n,B}^*$.

STUDENT. BOOT. CONF. INT.

$$\tilde{R}_{n,b}^* = \frac{\hat{\sigma}_{n,b}^* - \hat{\sigma}_n}{\hat{\sigma}_{n,b}^*},$$

$$\text{WHERE } \hat{\sigma}_{n,b}^{*2} = \frac{1}{(\bar{X}_{n,b}^*)^2} \frac{1}{n-1} \sum_{i=1}^n (Y_{i,b}^* - \hat{\sigma}_{n,b}^* X_{i,b}^*)^2$$

$$\left(\hat{\sigma}_n - \frac{\tilde{m}_{n,B}^*(1-\frac{\alpha}{2}) \hat{\sigma}_n}{\sqrt{n}}, \hat{\sigma}_n - \frac{\tilde{m}_{n,B}^*(\frac{\alpha}{2}) \hat{\sigma}_n}{\sqrt{n}} \right),$$

WHERE $\tilde{m}_{n,B}^*(\alpha)$ IS THE "SAMPLE" α -QUANTILE
FROM THE "BOOT. SAMPLE" $\tilde{R}_{n,1}^*, \dots, \tilde{R}_{n,B}^*$.

Ex 62) $X_1, \dots, X_n$ i.i.d. $U(0, \theta)$

$$\hat{\theta}_n = \max_{1 \leq i \leq n} \{X_i\}$$

$$\sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow[n \rightarrow \infty]{d} Z, \text{ WHERE } P(Z \leq x) = \begin{cases} e^{\frac{x}{\theta}}, & x < 0 \\ 1, & x \geq 0 \end{cases}$$

LET $R_n^* = \sqrt{n}(\hat{\theta}_n^* - \hat{\theta}_n)$,

WHERE $\hat{\theta}_n^* = \max_{1 \leq i \leq n} \{X_{i,b}^*\}$

NONPARAM. BOOT:

$$X_{1,b}^*, \dots, X_{n,b}^* \text{ i.i.d from } F_n(x) = \frac{1}{n} \sum_{i=1}^n \mathbb{1}\{X_i \leq x\}$$

PARAM. BOOT:

$$X_{1,b}^*, \dots, X_{n,b}^* \text{ i.i.d from } U(0, \hat{\theta}_n)$$

(Ex.)

$X_1, \dots, X_n$ i.i.d

$$F_X(x|\lambda) = \begin{cases} 1 - e^{-\lambda x}, & x \geq 0 \\ 0, & x < 0 \end{cases}$$

$H_0: \lambda(X_i) \in \{ \text{Exp}(\lambda) : \lambda > 0 \}$ $H_1: \text{not } H_0$

$$KS_n = \sup_{x \in \mathbb{R}} | \hat{F}_n(x) - F(x; \hat{\lambda}_n) |$$

$$\text{WHERE } \hat{F}_n(x) = \frac{1}{n} \sum_{i=1}^n \mathbb{1}\{X_i \leq x\}$$

$$F(x; \hat{\lambda}_n) = 1 - e^{-\hat{\lambda}_n x}, \quad \hat{\lambda}_n = \frac{1}{\bar{X}_n}$$

$$KS_{n,b}^* = \sup_{x \in \mathbb{R}} | \hat{F}_{n,b}^*(x) - F(x; \hat{\lambda}_{n,b}^*) |$$

$\uparrow$ $\text{ecdf } X_{1,b}^*, \dots, X_{n,b}^*$ $\uparrow$ $\frac{1}{\bar{X}_{n,b}^*}$

P-VALUE:

$$\frac{1 + \sum_{b=1}^B \mathbb{1}\{KS_{n,b}^* \geq KS_n\}}{B + 1}$$

$$CM_n = \frac{1}{n} \sum_{i=1}^n \left(\hat{F}_n(X_i) - F(X_i; \hat{\lambda}_n) \right)^2$$

$$AD_n = \frac{1}{n} \sum_{i=1}^n \frac{\left(\hat{F}_n(X_i) - F(X_i; \hat{\lambda}_n) \right)^2}{F(X_i; \hat{\lambda}_n) (1 - F(X_i; \hat{\lambda}_n))}$$

AVERAGE MARK 7TH GRADE

(E)

"COMPARING EXPECTATIONS IN TWO-SAMPLE PROBLEMS"

BOYS

$X_1, \dots, X_{n_1}$ i.i.d

INDEPENDENT

$$H_0: EX_1 = EY_1$$

GIRLS

$Y_1, \dots, Y_{n_2}$ i.i.d

$$H_1: EX_1 \neq EY_1$$

$$1. T_n = \frac{\bar{X}_{n_1} - \bar{Y}_{n_2}}{\sqrt{S^2_{n_1, n_2} \left(\frac{1}{n_1} + \frac{1}{n_2} \right)}}, \quad S^2_{n_1, n_2} = \frac{1}{n_1 + n_2 - 2} [(n_1 - 1)S_X^2 + (n_2 - 1)S_Y^2]$$

$$2. \tilde{T}_n = \frac{\bar{X}_{n_1} - \bar{Y}_{n_2}}{\sqrt{\frac{S_X^2}{n_1} + \frac{S_Y^2}{n_2}}}$$

3. NONPARAMETRIC BOOTSTRAP

$$\tilde{T}_{n,b}^* = \frac{\bar{X}_{n,b}^* - \bar{Y}_{n,b}^*}{\sqrt{\frac{S_{X,b}^{2*}}{n_1} + \frac{S_{Y,b}^{2*}}{n_2}}}$$

$X_{1,b}^*, \dots, X_{n_1,b}^*$

SAMPLE WITH REPLACEMENT FROM

$X_1 - \bar{X}_{n_1}, \dots, X_{n_1} - \bar{X}_{n_1}$

$Y_{1,b}^*, \dots, Y_{n_2,b}^*$

$Y_1 - \bar{Y}_{n_2}, \dots, Y_{n_2} - \bar{Y}_{n_2}$

4. PERMUTATION TEST

$$n = n_1 + n_2$$

- DENOTE: $\mathcal{Z} = (Z_1, \dots, Z_n)^T = (X_1, \dots, X_{n_1}, Y_1, \dots, Y_{n_2})^T$

$\rightarrow (Z_{1,b}^*, \dots, Z_{n,b}^*)^T$ BE A RANDOM PERMUTATION OF $\mathcal{Z}$

$$(X_{1,b}^*, \dots, X_{n_1,b}^*) = (Z_{1,b}^*, \dots, Z_{n_1,b}^*)^T$$

$$(Y_{1,b}^*, \dots, Y_{n_2,b}^*) = (Z_{n_1+1,b}^*, \dots, Z_{n_1+n_2,b}^*)^T$$

(Ex 70)

TWO-SAMPLE KOLM.-SMIRNOV TEST

$$K_{n_1, n_2} = \sup_x \left| \underset{\substack{\uparrow \\ \text{cdf from } X_1, \dots, X_{n_1}}} F_{n_1}(x) - \underset{\substack{\uparrow \\ \text{cdf } Y_1, \dots, Y_{n_2}}} G_{n_2}(x) \right|$$

$$K_{n_1, n_2, b}^* = \sup_x \left| \underset{\substack{\uparrow \\ \text{cdf } Z_{n_1, b}^*, \dots, Z_{n_1, b}^*}} F_{n_1, b}^*(x) - \underset{\substack{\uparrow \\ \text{cdf } Z_{n_1+n_2, b}^*, \dots, Z_{n_1+n_2, b}^*}} G_{n_2, b}^*(x) \right|$$

$$\text{P VALUE : } \frac{1 + \sum_{b=1}^B \mathbb{1}\{K_{n_1, n_2, b}^* \geq K_{n_1, n_2}\}}{B + 1}$$

Ex 7.1

$$\begin{pmatrix} Z_1 \\ I_1 \end{pmatrix}, \dots, \begin{pmatrix} Z_N \\ I_N \end{pmatrix}$$

$\{1, 2, \dots, b\}$ (pointing to Z_N)
 $\{A, B, C\}$ (pointing to I_N)

$$n_{jk} = \sum_{i=1}^N \mathbb{1}\{Z_i = j, I_i = k\}$$

PERM. RESAMPLE:

$$\begin{pmatrix} Z_{1,b}^* \\ I_1 \end{pmatrix}, \dots, \begin{pmatrix} Z_{N,b}^* \\ I_N \end{pmatrix}, \text{ WHERE}$$

$(Z_{1,b}^*, \dots, Z_{N,b}^*)$ IS PERMUT. OF $(Z_1, \dots, Z_N)$

$$\chi^2 = \sum_{j=1}^J \sum_{k=1}^K \frac{\left(n_{jk} - \frac{n_{j+} n_{+k}}{N} \right)^2}{\frac{n_{j+} n_{+k}}{N}}$$

$$\sqrt{n} \left(\hat{\beta}_n(\tau) - \beta(\tau) \right) \xrightarrow[n \rightarrow \infty]{d} N_2 \left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, V \right), \quad V = \left(E \left[\underline{X} \underline{X}^T f_{Y|X}(\hat{F}_{Y|X}^{-1}(\tau)) \right] \right)^{-1} E \left[\underline{X} \underline{X}^T \left(E \left[\underline{X} \underline{X}^T f_{Y|X}(\hat{F}_{Y|X}^{-1}(\tau)) \right] \right)^{-1} \right]$$

QUANTILE REGRESSION AND BOOTSTRAP

$$\underset{\substack{\uparrow \\ \text{FOOD EXPEND.}}}{F_{Y|X}}^{-1}(\tau) = \beta_0(\tau) + \beta_1(\tau) \underset{\substack{\uparrow \\ \text{INCOME}}}{X}$$

FOR SIMPLICITY ONLY $\tau = 0.25$ AND $\tau = 0.75$ CONSIDERED

NONPAR. BOOTSTRAP : $\begin{pmatrix} X_{n,b}^* \\ Y_{n,b}^* \end{pmatrix}, \dots, \begin{pmatrix} X_{n,b}^* \\ Y_{n,b}^* \end{pmatrix}$ FROM $\begin{pmatrix} X_1 \\ Y_1 \end{pmatrix}, \dots, \begin{pmatrix} X_n \\ Y_n \end{pmatrix}$

SUMMARY (... , SE = "BOOT") :

$$\hat{V}_{n,B} = \frac{1}{B-1} \sum_{b=1}^B \left(\hat{\beta}_{n,b}^* - \bar{\beta}_{n,B}^* \right) \left(\hat{\beta}_{n,b}^* - \bar{\beta}_{n,B}^* \right)^T,$$

$$\text{WHERE } \bar{\beta}_{n,B}^* = \frac{1}{B} \sum_{b=1}^B \hat{\beta}_{n,b}^*$$

AS. NORM. + BOOTSTRAP ESTIM. OF VARIANCE

$$\left(\hat{\beta}_1 \pm n^{-1/2} \alpha_{1/2} \sqrt{\hat{V}_{22,B}} \right),$$

WHERE $\hat{V}_{22,B}$ IS THE SECOND
DIAGONAL ELEMENT OF $\hat{V}_{n,B}$