NMAI059 Probability and statistics 1 Class 3

Robert Šámal

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Overview

Discrete random variables

Examples of discrete r.v.'s

Expectation

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Random variable

Often we are interested in a number given as a result of a random experiment.

- We throw a dart and measure the distance from the center of the dartboard.
- We roll a die until we get a six, then count how many rolls it took.
- In a quicksort algorithm (with a random choice of pivot) we measure the number of operations.

Definition

Given a probability space (Ω, \mathcal{F}, P) . We call a function $X : \Omega \to \mathbb{R}$ a discrete random variable, if Im(X) (range of X) is a countable set and if for every real x we have

$$\{\omega \in \Omega : X(\omega) = x\} \in \mathcal{F}.$$

PMF

Definition

Probability mass function, PMF of a discrete random variable X is a function $p_X : \mathbb{R} \to [0, 1]$ such that

$$p_X(x) = P(X = x) = P(\{\omega \in \Omega : X(\omega) = x\})$$

$$\blacktriangleright \sum_{x \in Im(X)} p_X(x) = ?$$

• S := Im(X) $Q(A) := \sum_{x \in A} p_X(x)$ $(S, \mathcal{P}(S), Q)$ is a discrete probability space.

For S = {s_i : i ∈ I} countable set of reals and c_i ∈ [0, 1] satisfying ∑_{i∈I} c_i = 1 there is a probability space and a discrete r.v. X on it such that p_X(s_i) = c_i for i ∈ I.

Another description – CDF

Definition

Cumulative distribution function, CDF of a r.v. X is a function

$$F_X(x) := P(X \le x) = P(\{\omega \in \Omega : X(\omega) \le x).$$

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• F_X is a nondecreasing function

$$\blacktriangleright \lim_{x \to -\infty} F_X(x) = 0$$

- $\blacktriangleright \lim_{x \to +\infty} F_X(x) = 1$
- \blacktriangleright F_X is right-continuous

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Expectation

Bernoulli/alternate distribution

- X = number of tails in one toss of a coin (not necessary a fair one)
- We write $X \sim Bern(p)$. (Sometimes Alt(p).)
- Given $p \in [0, 1]$.
- $\blacktriangleright p_X(1) = p$

▶
$$p_X(0) = 1 - p$$

- ▶ $p_X(k) = 0$ for $k \neq 0, 1$
- For an event $A \in \mathcal{F}$ we define *indicator random variable* I_A :

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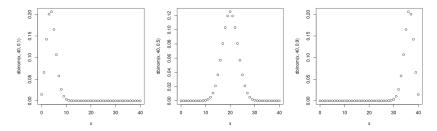
- $I_A(\omega) = 1$ if $\omega \in A$, $I_A(\omega) = 0$ otherwise.
- ► $I_A \sim Bern(P(A))$

Binomial distribution

- X = number of tails in n independent tosses of a loaded coin.
- Given $p \in [0, 1]$.
- We write $X \sim Bin(n, p)$.

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Binomial distribution: PMF



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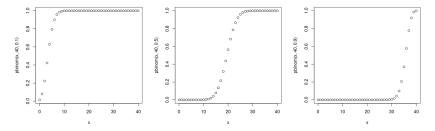
```
x <- 0:40

plot (x, dbinom(x,40,0.1))

plot (x, dbinom(x,40,0.5))

plot (x, dbinom(x,40,0.9))
```

Binomial distribution: CDF



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```
x <- 0:40

plot (x, pbinom(x,40,0.1))

plot (x, pbinom(x,40,0.5))

plot (x, pbinom(x,40,0.9))
```

Hypergeometric distribution

► X = the number of red balls we get out of n, when the urn contains K red out of N balls

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- ▶ Given *n*, *N*, *K*.
- We write $X \sim Hyper(N, K, n)$.

•
$$p_X(k) = P(X = k) = \frac{\binom{K}{k}\binom{N-K}{n-k}}{\binom{N}{n}}$$

Poisson distribution

• We write
$$X \sim Pois(\lambda)$$
.

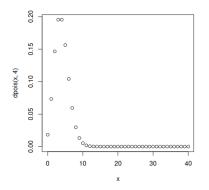
• Given real
$$\lambda > 0$$
.

$$\blacktriangleright p_X(k) = \frac{\lambda^k}{k!} e^{-\lambda}$$

•
$$Pois(\lambda)$$
 is a limit of $Bin(n, \lambda/n)$

► X describes, e.g., the number of emails we get in a day.

Poisson distribution: PMF



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```
x <- seq(0,40,by=1)
plot(x,dpois(x,4))
```

Poisson paradigm

A₁,..., A_n are (almost-)independent events with P(A_i) = p_i, λ = ∑_i p_i. Suppose n is large, each of p_i small. Then it is approximately true that

$$\sum_{i=1}^{n} I_{A_i} \sim Pois(\lambda).$$

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Geometric distribution

- X = number of coin tosses till we get a tail
- We write $X \sim Geom(p)$.

• Given
$$p \in [0, 1]$$
.

- $p_X(k) = (1-p)^{k-1}p$, for k = 1, 2, ...
- Some people call this distribution shifted geometric, the normal geometric would then be distribution of *X* − 1, that is the number of unsuccessful tosses.

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Expectation

Definition

Given a discrete r.v. X, its expectation is denoted by $\mathbb{E}(X)$ and defined by

$$\mathbb{E}(X) = \sum_{x \in Im(X)} x \cdot P(X = x),$$

whenever the sum is defined.

Suppose X is defined on a discrete space (Ω, F, P). Then we can also define the expectation by the following formula:

$$\mathbb{E}(X) = \sum_{\omega \in \Omega} X(\omega) P(\{\omega\}).$$

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Law Of The Unconscious Statistician

► For a real function g and a discrete r.v. X, the function Y = g(X) is also a discrete r.v.

Theorem (LOTUS)

For a real function g and a discrete r.v. X, we have

$$\mathbb{E}(g(X)) = \sum_{x \in Im(X)} g(x)P(X = x)$$

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whenever the sum is defined.

Properties of ${\ensuremath{\mathbb E}}$

Theorem

Suppose X, Y are discrete r.v. and $a, b \in \mathbb{R}$.

1. If $P(X \ge 0) = 1$ and $\mathbb{E}(X) = 0$, then P(X = 0) = 1.

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- **2.** If $\mathbb{E}(X) \ge 0$ then $P(X \ge 0) > 0$.
- **3**. $\mathbb{E}(a \cdot X + b) = a \cdot \mathbb{E}(X) + b$.

4.
$$\mathbb{E}(X+Y) = \mathbb{E}(X) + \mathbb{E}(Y).$$

Variance

Definition

Variance of a r.v. X is the number $\mathbb{E}((X - \mathbb{E}X)^2)$. It is denoted by var(X).

Theorem

$$var(X) = \mathbb{E}(X^2) - \mathbb{E}(X)^2$$

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Conditional expectation

Definition

Let X be a discrete r.v. and P(B) > 0. Conditional expectation of X given B is

$$\mathbb{E}(X \mid B) = \sum_{x \in Im(X)} x \cdot P(X = x \mid B),$$

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whenever the sum is defined.

Law Of Total Expectation

Theorem Suppose B_1, B_2, \ldots is a partition of Ω and $A \in \mathcal{F}$. Then

$$\mathbb{E}(X) = \sum_{i} \mathbb{E}(X \mid B_i) P(B_i),$$

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whenever the sum is defined. (Terms with $P(B_i) = 0$ are counted as 0.)

Law Of Total Expectation

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